

10.13. Test for homogeneity in populations

Example. Humans red ~~test~~ blood cells can be divided into 4 hereditary blood-types: A, B, AB and O. An international health organization wishes to find out if the relative occurrence is the same in 3 particular ethnic groups: G1, G2 and G3. Random samples of sizes 400, 200 and 100 persons are drawn from the 3 groups and their blood type is decided for each person.

Contingency Table

Blood type \ Group		G1	G2	G3	Total
B1	A	176 (192)	112 (96)	48 (48)	336
B2	B	41 (37.14)	16 (18.57)	8 (9.39)	65
B3	AB	19 (16.57)	6 (8.29)	4 (4.14)	29
B4	O	164 (154.29)	66 (77.14)	40 (38.57)	270
Total		400	200	100	700

Define $p_{ij} = P(BI \cap GJ)$, $i = 1, 2, 3, 4$, $j = 1, 2, 3$

Homogeneity means $p_{ij} = p_i$, $j = 1, 2, 3$ for all i

$$\hat{p}_1 = \frac{336}{700}, \hat{p}_2 = \frac{65}{700}, \hat{p}_3 = \frac{29}{700} \text{ and } \hat{p}_4 = \frac{270}{700} = 1 - \hat{p}_1 - \hat{p}_2 - \hat{p}_3$$

let m_{ij} be ^{estimated} expected numbers in cell ij

$$m_{ij} = n_j \cdot \frac{m_i}{700} = n_j \cdot \hat{p}_i$$



total for column j

$$H_0 : p_{i1} = p_{i2} = p_{i3} = p_i \quad i = 1, 2, \dots, 4$$

$$\text{Degrees of freedom: } (r-1)c - (r-1) = (r-1)(c-1) = 3 \cdot 2 = 6$$

$$\begin{aligned} \chi^2 = & \frac{(176-192)^2}{192} + \frac{(41-37.14)^2}{37.14} + \frac{(19-16.57)^2}{16.57} + \frac{(164-154.29)^2}{154.29} \\ & + \frac{(112-96)^2}{96} + \frac{(16-18.57)^2}{18.57} + \frac{(6-8.29)^2}{8.29} + \frac{(66-77.14)^2}{77.14} \\ & + \frac{(48-48)^2}{48} + \frac{(8-9.29)^2}{9.29} + \frac{(4-4.14)^2}{4.14} + \frac{(40-38.57)^2}{38.57} = 8.2 \end{aligned}$$

$$\chi^2_{0.05}(6) = 12.59 \Rightarrow \text{no reason to reject } H_0 \text{ on a}$$

5% level.

Nonparametric Statistics

Nonparametric tests

Tests where no or very little knowledge of the distributions are assumed. These are often formulated on the median, $\tilde{\mu}$, instead of the mean.

Note. The t -test and the F -test is robust to slight deviations from normality.

The sign test (Assumes only continuous distributions)

Example

A rectangle where the ratio between the height and the length is 0.618 (the golden ratio) are called ^agolden rectangles. Did the Shoshoni indians know about this?

Below is given 8 such ratios found on leather goods made of Shoshoni indians.

X_i : 0.693 0.662 0.690 0.606 0.570 0.749 0.672 0.628

Want to test the following hypothesis

$$H_0: \tilde{\mu} = 0.618$$

$$H_1: \tilde{\mu} \neq 0.618$$

X_i	$X_i - 0.618$	signs of difference
0.693	0.075	+
0.662	0.044	+
0.690	0.072	+
0.606	-0.012	-
0.570	-0.048	-
0.749	0.131	+
0.672	0.054	+
0.628	0.010	+

Under H_0 : $P(X < 0.618) = P(X > 0.618) = 0.5$

Let Y be the number of rectangles with ratio less than 0.618. $Y \sim B(8, 0.5)$. $Y_{obs} = 2$.

The p-value is $2 P(Y \leq 2) = \underline{0.289}$

Assume the distribution is continuous and symmetric

Then we can use the Wilcoxon signed-rank test.

We then rank the observations ~~x_i~~ $(x_i - \tilde{\mu}_0)$ according to their absolute values.

We get:

	0.008	0.012	0.044	0.048	0.054	0.072	0.075	0.131
Rank:	1	2	3	4	5	6	7	8
Sign:	+	-	+	-	+	+	+	+

let W_+ be the rank of those with a positive sign and

W_- — — — — — negative sign

and let W be the smallest of W_+ and W_- .

For testing.

H_0	H_1
$\tilde{\mu} = \tilde{\mu}_0$	$\tilde{\mu} < \tilde{\mu}_0$ use W_+
	$\tilde{\mu} > \tilde{\mu}_0$ use W_-
	$\tilde{\mu} \neq \tilde{\mu}_0$ use W

In this case $W_- = 6$ and $W_+ = 30 \Rightarrow W = W_-$

There are ~~2^8~~ ^{$(2^8 = 256)$} possible ways to place out signs

$$\begin{aligned}
 P(W_- = 0) &= \frac{1}{2^8} & P(W_- = 3) &= \frac{2}{2^8} & P(W_- = 6) &= \frac{3}{2^8} \\
 P(W_- = 1) &= \frac{1}{2^8} & P(W_- = 4) &= \frac{2}{2^8} & \Rightarrow P(W_- \leq 6) &= \frac{13}{2^8} = \frac{13}{256} \\
 P(W_- = 2) &= \frac{1}{2^8} & P(W_- = 5) &= \frac{3}{2^8}
 \end{aligned}$$

For a two sided test. this gives a p-value of $\frac{26}{252} > 0.05$

$\Rightarrow$ no reason to reject H_0 at a 5% level.

In order to avoid a lot of tedious calculations the critical values are given in Table A.16.

It says that H_0 should be rejected i.f. W (or W_-) ≤ 4 .

Approximation to the normal distribution

Now. let $U = W_+ - W_-$

and define $I_j = \begin{cases} 1 & \text{i.f. } x_j - \tilde{\mu}_0 > 0 \\ -1 & \text{i.f. } x_j - \tilde{\mu}_0 < 0 \end{cases}$

Then $U = \sum_{j=1}^m R_j I_j = \sum_{j=1}^m j I_j$ in distribution.
Rank of $\{x_j - \tilde{\mu}_0\}$

Under H_0 $P(I_j = 1) = P(I_j = -1) = \frac{1}{2}$ and we get.

$$E[I_j] = 0, \quad \text{Var}[I_j] = 1$$

$$\text{and } E[U] = 0 \quad \text{and } \text{Var}[U] = \sum_{j=1}^m j^2 \cdot 1 = \frac{n(n+1)(2n+1)}{6}.$$

$$W_- + W_+ = \sum_{j=1}^m j = \frac{n(n+1)}{2} \quad \text{and since } W_- = W_+ - U$$

$$\text{we get } 2W_+ = U + \frac{n(n+1)}{2}$$

$$\text{Therefore } E[W_+] = \frac{n(n+1)}{4} \quad \text{and } \text{Var}[W_+] = \frac{n(n+1)(2n+1)}{24}.$$

For $n \geq 15$ the ~~distribution~~ of W_+ (and W_-) $\approx N\left(\frac{n(n+1)}{4}, \frac{n(n+1)(2n+1)}{24}\right)$.